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# GENERATING FUNCTIONS

**Def.4.1:** Given a collection of real numbers  $(a_k)_{k\geq 0}$ , the function

$$G(s)=G_a(s)\stackrel{\mathsf{def}}{=}\sum_{k=0}^\infty a_k\,s^k$$

is called the generating function of  $(a_k)_{k\geq 0}$ .

If  $G_a(s)$  is analytic near the origin, then

$$a_k = rac{1}{k!} \, rac{d^k}{ds^k} \, G_{\mathsf{a}}(s) igert_{s=0} \; .$$

This result is often referred to as the **uniqueness** property of generating functions.

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**Def.4.2:** If X is a discrete random variable with values in  $\mathbb{Z}^+ \stackrel{\text{def}}{=} \{0, 1, ...\}$ , its (*probability*) generating function,

$$G(s) \equiv G_X(s) \stackrel{\text{def}}{=} \mathsf{E}(s^X) = \sum_{k=0}^{\infty} s^k \mathsf{P}(X=k), \qquad (1.1)$$

is just the generating function of the pmf  $\{p_k\} \equiv \{P(X = k)\}$  of X.

Recall that the moment generating function  $M_X(t) \stackrel{\text{def}}{=} \mathsf{E}(e^{tX})$  of a random variable X is just  $\sum_{k\geq 0} \frac{\mathsf{E}(X^k)}{k!} t^k$ .

Why do we introduce both  $G_X(s)$  and  $M_X(t)$ ?



**Theorem 4.3:** If X and Y are independent random variables with values in  $\{0, 1, 2, ...\}$  and  $Z \stackrel{\text{def}}{=} X + Y$ , then their generating functions satisfy

$$G_Z(s) = G_{X+Y}(s) = G_X(s) G_Y(s).$$

<u>Recall</u>: if X and Y are discrete random variables, and f,  $g : \mathbb{Z}^+ \to \mathbb{R}$  are arbitrary functions, then f(X) and g(Y) are independent random variables and  $E[f(X)g(Y)] = Ef(X) \cdot Eg(Y)$ .

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**Example 4.4:** If  $X_1, X_2, ..., X_n$  are *i.i.d.r.v.* with values in  $\{0, 1, 2, ...\}$  and if  $S_n = X_1 + \cdots + X_n$ , then

$$G_{S_n}(s) = G_{X_1}(s) \dots G_{X_n}(s) \equiv \left[G_X(s)\right]^n$$
.

**Example 4.5:** Let  $X_1, X_2, ..., X_n$  be i.i.d.r.v. with values in  $\{0, 1, 2, ...\}$  and let  $N \ge 0$  be an integer-valued random variable independent of  $\{X_k\}_{k\ge 1}$ . Then  $S_N \stackrel{\text{def}}{=} X_1 + \cdots + X_N$  has generating function

$$G_{S_N}(s) = G_N(G_X(s)).$$

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**Example 4.6:** [Renewals] Imagine a diligent janitor who replaces a light bulb the same day as it burns out. Suppose the first bulb is put in on day 0 and let  $X_i$  be the lifetime of the ith light bulb. Let the individual lifetimes  $X_i$  be i.i.d.r.v.'s with values in  $\{1, 2, ...\}$  and have a common distribution with generating function  $G_f(s)$ . Define  $r_n \stackrel{\text{def}}{=} P(a \text{ light bulb was replaced on day } n)$  and  $f_k \stackrel{\text{def}}{=} P(\text{the first light bulb was replaced on day } k)$ . Then  $r_0 = 1$ ,  $f_0 = 0$ , and for  $n \ge 1$ ,

$$r_n = f_1 r_{n-1} + f_2 r_{n-2} + \cdots + f_n r_0 = \sum_{k=1}^n f_k r_{n-k}.$$

A standard computation implies that  $G_r(s) = 1 + G_f(s) G_r(s)$  for all  $|s| \le 1$ , so that  $G_r(s) = 1/(1 - G_f(s))$ .

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In general, we say a sequence  $(c_n)_{n\geq 0}$  is the convolution of  $(a_k)_{k\geq 0}$ and  $(b_m)_{m\geq 0}$  (write  $c = a \star b$ ), if

$$c_n = \sum_{k=0}^n a_k b_{n-k}, \qquad n \ge 0.$$

**Theorem 4.7:** [Convolution thm] If  $c = a \star b$ , then the generating functions  $G_c(s)$ ,  $G_a(s)$ , and  $G_b(s)$  satisfy

$$G_c(s) = G_a(s) G_b(s)$$
.

**Example 4.8:** Let  $X \sim \text{Poi}(\lambda)$  and  $Y \sim \text{Poi}(\mu)$  be independent. Then Z = X + Y is  $\text{Poi}(\lambda + \mu)$ .

Solution. A straightforward computation gives  $G_X(s) = e^{\lambda(s-1)}$ , and Theorem 4.3 implies  $G_Z(s) = G_X(s) G_Y(s) \equiv e^{(\lambda+\mu)(s-1)}$ , so that the result follows by uniqueness.

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Notice: if  $X_1, X_2, ..., X_n$  are independent Bernoulli rv's with parameter  $p \in [0, 1]$ , each  $X_k$  has the generating function  $G(s) = qs^0 + ps^1 = q + ps$ . If  $S_n = \sum_{k=1}^n X_k$ , then

$$G_{S_n}(s) = G_{X_1}(s) G_{X_2}(s) \dots G_{X_n}(s) = (G(s))^n = (q + ps)^n$$

By the Binomial theorem, this reads

$$G_{S_n}(s) = \sum_{k=0}^n \binom{n}{k} p^k q^{n-k} s^k,$$

ie.,  $S_n \sim Bin(n, p)$ .

**Example 4.9:** If  $X \sim Bin(n, p)$  and  $Y \sim Bin(m, p)$  are independent, then  $X + Y \sim Bin(n + m, p)$ .

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### Computation of moments

**Theorem 4.10:** If a rv X has generating function G(s), then a) E(X) = G'(1).

b) more generally,  $\mathsf{E}ig[X(X-1)\dots(X-k+1)ig]=G^{(k)}(1);$ 

here,  $G^{(k)}(1)$  is the shorthand for  $\lim_{s\uparrow 1} G^{(k)}(s)$ , the limiting value of the kth left derivative of G(s) at s = 1.

The quantity E[X(X-1)...(X-k+1)] is known as the *k*th factorial moment of *X*.

It is a straightforward exercise to show that

$$\operatorname{Var}(X) = G_X''(1) + G_X'(1) - \left(G_X'(1)
ight)^2.$$

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**Proof:** b) Fix  $s \in (0, 1)$  and differentiate G(s) k times to get $G^{(k)}(s) = \mathbb{E}[s^{X-k}X(X-1)\dots(X-k+1)].$ 

Taking the limit  $s \uparrow 1$  and using the Abel theorem, we obtain

$$G^{(k)}(s) 
ightarrow \mathsf{E}ig[X(X-1)\dots(X-k+1)ig]$$
 .

Notice also that

$$\lim_{s \nearrow 1} G_X(s) \equiv \lim_{s \nearrow 1} \mathsf{E}[s^X] = \mathsf{P}(X < \infty).$$

This allows us to check whether a variable is finite, if we do not know this *apriori*.

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#### RECURRENCES

Generating functions are very useful in solving **recurrences**, especially when combined with the following algebraic fact.

**Lemma 4.12:** Let f(x) = g(x)/h(x) be a ratio of two polynomials without common roots. Let  $\deg(g) < \deg(h) = m$  and suppose that the roots  $a_1, \ldots, a_m$  of h(x) are all distinct. Then f(x) can be decomposed into a sum of partial fractions, i.e., for some constants  $b_1, b_2, \ldots, b_m$ ,

$$f(x) = \frac{b_1}{a_1 - x} + \frac{b_2}{a_2 - x} + \dots + \frac{b_m}{a_m - x}$$

Since

$$\frac{b}{a-x} = \frac{b}{a} \sum_{k \ge 0} \left(\frac{x}{a}\right)^k = \sum_{k \ge 0} \frac{b}{a^{k+1}} x^k \,,$$

a partial fraction sum can be easily written as a power series.

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**Example 4.13:** Let  $a_n$  be the probability that n independent Bernoulli trials (with success probability p) result in an **even number of successes**. Find the generating function of  $a_n$ .

**Solution.** The event under consideration occurs if an **initial failure** at the first trial is followed by an even number of successes or if an **initial success** is followed by an odd number of successes. Therefore,  $a_0 = 1$  and

$$a_n = (1-p) a_{n-1} + p (1-a_{n-1}), \qquad n \ge 1.$$

Multiplying these equalities by  $s^n$  and adding we get (with q = 1 - p)

$$G_a(s) - 1 = qs \ G_a(s) + p \sum_{n \ge 1} s^n - ps \ G_a(s) = (q - p)s \ G_a(s) + rac{ps}{1 - s} \, ,$$

and after rearranging,

$$G_a(s) = \left(1 + rac{ps}{1-s}\right) / \left(1 - (q-p)s\right) = rac{1}{2} \left(rac{1}{1-s} + rac{1}{1-(q-p)s}\right).$$

As a result,  $a_n = \left(1 + (q-p)^n\right)/2.$ 

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**Example 4.14:** A biased coin is tossed repeatedly; on each toss, it shows a "head" with probability p. Let  $r_n$  be the probability that a sequence of n tosses never has two "heads" in a row. Show that  $r_0 = 1$ ,  $r_1 = 1$ , and for all n > 1,  $r_n = qr_{n-1} + pqr_{n-2}$ , where q = 1 - p. Find the generating function of the sequence  $(r_n)_{n \ge 0}$ .

**Solution.** Every sequence of *n* tosses contributing to  $r_n$  starts either with T or with HT; therefore, for all n > 1,  $r_n = qr_{n-1} + pqr_{n-2}$  (where q = 1 - p). Multiplying these equalities by  $s^n$  and summing, we get

$$G_r(s) = \sum_{n \ge 0} r_n s^n = 1 + s + qs \sum_{n \ge 2} r_{n-1} s^{n-1} + pqs^2 \sum_{n \ge 2} r_{n-2} s^{n-2}$$
$$= (qs + pqs^2)G_r(s) + 1 + ps,$$

so that

$$\mathit{G_r}(s) = rac{1+
ho s}{1-qs-
ho qs^2}\,.$$

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### CONTINUITY THEOREM

**Theorem 4.15:** Let for every fixed *n* the sequence  $a_{0,n}$ ,  $a_{1,n}$ , ... be a **probability distribution**, i.e.,  $a_{k,n} \ge 0$  and  $\sum_{k\ge 0} a_{k,n} = 1$ , and let  $G_n(s) = \sum_{k\ge 0} a_{k,n} s^k$  be the corresponding generating function. Then

$$^orall k \geq 0 \quad a_{k,n} o a_k \quad \Longleftrightarrow \quad ^orall s \in [0,1) \quad \mathcal{G}_n(s) o \mathcal{G}(s) \, ,$$

where  $G(s) = \sum_{k\geq 0} a_k s^k$ , the generating function of the limiting sequence  $(a_k)_{k\geq 0}$ .

The convergence  $\{a_{k,n}\}_{k\geq 0} \rightarrow \{a_k\}_{k\geq 0}$  is known as the **convergence in distribution**!



**Example 4.16:** If  $X_n \sim Bin(n, p)$  with  $p = p_n$  satisfying  $n \cdot p_n \rightarrow \lambda$  as  $n \rightarrow \infty$ , then

$$G_{X_n}(s) \equiv \left(1 + p_n(s-1)\right)^n \to \exp\{\lambda(s-1)\},$$

so that the distribution of  $X_n$  converges to that of  $X \sim Poi(\lambda)$ .

R script!

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# GENERATING FUNCTIONS

By the end of this section you should be able to:

- define ordinary, probability, and moment generating functions;
- derive the value of the *n*th term of a sequence from the corresponding generating function;
- state and apply the theorem about generating functions of convolutions;
- use probability generating functions to compute moments of random variables;
- state and apply the uniqueness and continuity theorems for generating functions;
- use generating functions in solving recurrent relations;
- explain the role of generating functions in proving convergence in distribution.

# Problem GF-17

Two players play a game called *heads or tails*. In this game, a coin coming up heads with probability p is tossed consecutively. Each time a head comes up Player I wins 1 pence, otherwise she loses 1 pence. Let  $X_k \in \{-1, 1\}$  denote the outcome of the *k*th trial, and let  $S_n$ ,  $n \ge 0$ , be the total gain of Player I after n trials,  $S_n = X_1 + \cdots + X_n$ , where different outcomes are assumed independent. Let T be the moment when Player I is first in the lead, ie.,

 $S_k \leq 0$  for k < T, and  $S_T = 1$ .

(A) find the generating function  $G_T(s) = E(s^T)$  of T;

(B) for which values of p is T a proper random variable, i.e., when  $P(T < \infty) \equiv \lim_{s \neq 1} G_T(s) = 1$ ?

(C) compute the expectation  $\mathsf{E}T$  of T.

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## Solution

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Let  $f_{ij}^{(k)} = P(S_1 \neq j, ..., S_{k-1} \neq j, S_k = j | S_0 = i)$  be the probability to **first** hit state *j* after *k* jumps, starting from *i*. Then

$$\mathsf{P}(T=k) = f_{01}^{(k)}$$
 and  $G_T(s) = \sum_{k=0}^{\infty} f_{01}^{(k)} s^k$ .

On the event  $\{X_1 = 1\}$  we have T = 1, and on the event  $\{X_1 = -1\}$  we have  $T = 1 + T_2$ , where  $T_2$  is the time to **first** hit state 1 starting from -1. By the partition theorem,

$$G_T(s) = \mathsf{E}(s^T \mid X_1 = 1)\mathsf{P}(X_1 = 1) + \mathsf{E}(s^T \mid X_1 = -1)\mathsf{P}(X_1 = -1)$$
  
or (with  $q = 1 - p$ )

$$G_{T}(s) = ps + qs\mathsf{E}(s^{T_2}) = ps + qs G_{T_2}(s). \tag{*}$$

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To get from state -1 to state 1 we have to pass through 0, so that

$$\mathsf{P}(T = m) = f_{-1,1}^{(m)} = \sum_{\ell=1}^{m-1} f_{-1,0}^{(\ell)} f_{0,1}^{(m-\ell)}$$

where, by homogeneity,  $f_{-1,0}^{(\ell)} \equiv f_{0,1}^{(\ell)}$ ; the convolution theorem now implies  $G_{T_2}(s) = G_T(s)G_T(s) = (G_T(s))^2$ . Substituting this into (\*) and solving for  $G_T(s)$ , we get (since  $G_T(s) \to 0$  if  $s \to 0$ !)

$$G_T(s) = rac{1 - \sqrt{1 - 4 
ho q s^2}}{2 q s} = rac{2 
ho s}{1 + \sqrt{1 - 4 
ho q s^2}}$$

Now,

А

$$\mathsf{P}(T<\infty)=\mathcal{G}_{T}(1)=rac{1-|p-q|}{2q}=egin{cases}1\,,&p\geq q\,,\powerbox{}\p$$

Similarly (do the computation!),

$$\mathsf{E}T = G_T'(1) = \begin{cases} 1/(p-q)\,, & p > q\,,\\ +\infty\,, & p = q\,. \end{cases}$$
  
Iso, for  $p < q$  we have  $\mathsf{P}(T = \infty) = 1 - \frac{p}{q} > 0$ , so that  $\mathsf{E}T = \infty$ .