PROBABILITY II (MATH 2647)

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http://maths.dur.ac.uk/stats/courses/ProbMC2H/Probability2H.html

or via DUO

This term we shall consider:

- Review of Core A Probability
- Elements of convergence and integration (*)
- Generating functions
- Markov chains
- Random walks (*)
- ...

(*) important if you are going to do Math Finance!

Lectures: Tuesday 11, Wednesday 11 in CM101 weeks 1-10,20 Problems Classes: Thursday 12 in CM101 weeks 4,6,8,10 Tutorials: according to your timetable weeks 3,5,7,9,20

Exam: 2 hours, 6 questions (3+3)

"Credit will be given for the best TWO answers from Section A and the best TWO answers from Section B. Questions in Section B carry ONE and a HALF times as many marks as those in Section A. Use of electronic calculators is forbidden."

subject to confirmation for this year!

Homeworks: set on Wednesday, due in CM116 on Thursday the week after

exact deadline will be communicated via email!

the best preparation for the exam!

data!

All course materials available from the course webpage!

SAMPLE SPACE & EVENTS

Sample space Ω is a collection of all possible outcomes of a probabilistic experiment;

Event is a collection of possible outcomes, ie., a subset of the sample space.

E.g., the *impossible* event \emptyset , the *certain* event Ω ;

also, if $A \subset \Omega$ and $B \subset \Omega$ are events, one considers $A \cup B$ (**A** or **B**), $A \cap B$ (**A** and **B**), $A^{c} \equiv \Omega \setminus A$ (not **A**), $A \setminus B$ (**A** but not **B**).

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σ -fields

Definition

Let \mathcal{F} be a collection of subsets of Ω . We shall call \mathcal{F} a σ -field if it has the following properties:

1.
$$\emptyset \in \mathcal{F}$$
;
2. if $A_1, A_2, \dots \in \mathcal{F}$, then $\bigcup_{k=1}^{\infty} A_k \in \mathcal{F}$;
3. if $A \in \mathcal{F}$ then $A^c \in \mathcal{F}$

Remark If condition 2. is replaced with

2'. if $A_1, A_2 \in \mathcal{F}$, then $A_1 \cup A_2 \in \mathcal{F}$; equivalently, with

2". if $A_1, A_2, \ldots, A_m \in \mathcal{F}$, then $\bigcup_{k=1}^m A_k \in \mathcal{F}$; the collection \mathcal{F} is called a **field**.

Remark

Condition 2. in the definition above is important: indeed, in a standard coin-flipping experiment let

$$A_k = \{$$
first 'head' occurs on kth flip $\}$;

then

$$\left\{ \mathsf{a} \text{ 'head' observed} \right\} = \bigcup_{k=1}^{\infty} A_k \,.$$

This also explains why one needs to study infinite sample spaces.

PROBABILITY DISTRIBUTION

Definition

Let Ω be a sample space, and \mathcal{F} be a σ -field of events in Ω . A **probability distribution** P on (Ω, \mathcal{F}) is a collection of *numbers* P(A), $A \in \mathcal{F}$, possessing the following properties:

- A1 for every event $A \in \mathcal{F}$, $\mathsf{P}(A) \ge 0$;
- A2 P(Ω) = 1;
- A3 for any pair of *incompatible* events A and B, $P(A \cup B) = P(A) + P(B);$
- A4 for any *countable* collection A_1, A_2, \ldots of *mutually incompatible* events,

$$\mathsf{P}\Big(\bigcup_{k=1}^{\infty}A_k\Big)=\sum_{k=1}^{\infty}\mathsf{P}\big(A_k\big).$$

recall: events A, B are incompatible if $A \cap B = \emptyset$.

PROBABILITY SPACE

Definition

A probability space is a triple $(\Omega, \mathcal{F}, \mathsf{P})$, where Ω is a sample space, \mathcal{F} is a σ -field of events in Ω , and $\mathsf{P}(\cdot)$ is a probability measure on (Ω, \mathcal{F}) .

In what follows we shall **always assume** that some probability space $(\Omega, \mathcal{F}, \mathsf{P})$ is fixed.

Why do we care?

- To avoid pathological events which cannot be assigned a probability
 Do not worry, no such problems in this course!
- To avoid "paradoxes" (e.g., Bertrand's paradox)

• . . .

Bertrand's paradox

Suppose a chord of a circle is chosen at random. What is the probability that the chord is longer than a side of an equilateral triangle inscribed in the circle?

Different types of "randomness":

"random endpoints" / "random radius" / "random midpoint" give different results!

Some properties

The following properties are immediate from the above axioms: P1 for any pair of events A, B in Ω we have

$$P(B \setminus A) = P(B) - P(A \cap B),$$
$$P(A \cup B) = P(A) + P(B \setminus A);$$

in particular, $P(A^c) = 1 - P(A)$;

P2 if events A, B in Ω are such that $\varnothing \subseteq A \subseteq B \subseteq \Omega$, then

$$0 = \mathsf{P}(\emptyset) \le \mathsf{P}(A) \le \mathsf{P}(B) \le \mathsf{P}(\Omega) = 1$$
.

P3 if A_1, A_2, \ldots, A_n are events in Ω , then

$$\mathsf{P}\big(\bigcup_{k=1}^n A_k\big) \le \sum_{k=1}^n \mathsf{P}(A_k)$$

with the inequality becoming an **equality** if these events are **mutually incompatible**;

Some properties [cont'd]

Probability measure is continuous along monotone sequences:

If
$$A_1 \subseteq A_2 \subseteq ...$$
 are events and $A = \lim_{k \to \infty} A_k \equiv \bigcup_{k=1}^{\infty} A_k$, then

$$P(A) = \lim_{k \to \infty} P(A_k).$$

If
$$B_1 \supseteq B_2 \supseteq \ldots$$
 are events and $B = \lim_{k \to \infty} B_k \equiv \bigcap_{k=1}^{\infty} B_k$, then

$$P(B) = \lim_{k \to \infty} P(B_k).$$

Example:

A coin with success probability $p \in (0, 1)$ is tossed repeatedly. Let S_k be the event "no successes in the first k tosses". Then $S_k \to S$, where S is the event "no successes were observed".

As the sequence $\{S_k\}_{k\geq 1}$ is **monotone** with $P(S_k) = (1-p)^k$, we get $P(S) = \lim_{k \to \infty} P(S_k) = 0.$

INDEPENDENCE OF EVENTS

Definition

We say that events A and B are *independent* if

$$\mathsf{P}(A \cap B) = \mathsf{P}(A) \mathsf{P}(B) \,.$$

Definition

A collection A_1, A_2, \ldots of events is (mutually) *independent* if its **every finite** sub-collection is independent, i.e., for all $k \in \mathbb{N}$ and all $i_1, \ldots, i_k \geq 1$,

$$\mathsf{P}\Big(\bigcap_{\ell=1}^k A_{i_\ell}\Big) = \prod_{\ell=1}^k \mathsf{P}\big(A_{i_\ell}\big)\,.$$

Notion of *independence* is the key distinction between probability and (real) analysis (measure theory)!

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Example:

A coin showing heads with probability p is tossed repeatedly. Let R be the event "sub-word TT occurs before sub-word HH",

$$R = \left\{ TT, HTT, THTT, \ldots \right\}.$$

Assuming independence of individual outcomes, find P(R).

Consider $r_k = \{ "TT \text{ first occurs at kth toss and no } HH \text{ observed"} \}$. Then $P(r_k) = p^m q^{k-m}$ with q = 1 - p and $m = \lfloor k/2 \rfloor - 1$, implying $P(R) = (1+p)q^2/(1-pq)$.

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CONDITIONAL PROBABILITY

Definition

The *conditional probability* of event A given event B such that P(B) > 0, is

$$\mathsf{P}(A \mid B) \stackrel{\mathsf{def}}{=} \frac{\mathsf{P}(A \cap B)}{\mathsf{P}(B)}$$

It follows that

$$\mathsf{P}(A \cap B) = \mathsf{P}(A) \mathsf{P}(B | A) = \mathsf{P}(B) \mathsf{P}(A | B);$$

in particular, P(A | B) = P(A) if A and B are independent.

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LAW OF TOTAL PROBABILITY

If events B_1, \ldots, B_n form a *partition* of Ω , then for any event A,

$$\mathsf{P}(A) = \sum_{k=1}^{n} \mathsf{P}(A \cap B_k) = \sum_{k=1}^{n} \mathsf{P}(B_k) \cdot \mathsf{P}(A \mid B_k).$$

Example:

One independently rolls twenty standard six-sided dice. What is the probability that the sum of observed results is divisible by 6?

Example (Renewals)

A diligent janitor replaces a light bulb the instant it burns out. Suppose:

- that the first bulb is put in at time 0
- that the lifetimes of different light bulbs are independent
- and have the same distribution.

Let $A = \{a \text{ light bulb was replaced at time } n\}$, with $r_n = P(A)$; $B_k = \{\text{the first light bulb was replaced at time } k\}$, $f_k = P(B_k)$.

The law of total probability then implies

$$r_n = P(A) = \sum_{k=1}^n P(B_k) \cdot P(A | B_k) = \sum_{k=1}^n f_k r_{n-k}.$$

SEQUENTIAL BAYES FORMULA

If events A, B, C are such that $P(B \cap C) > 0$, then $P(A \cap B \mid C) = P(A \mid B \cap C) P(B \mid C)$.

Equivalently,

 $\mathsf{P}(A \cap B \cap C) = \mathsf{P}(A \mid B \cap C) \mathsf{P}(B \mid C) \mathsf{P}(C).$

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Example:

Consecutive winnings of a game, X_k , are iid random variables with

$$\mathsf{P}(X=1) = 1 - \mathsf{P}(X=-1) = p$$
 ,

so the total winning is $S_0 = 0$, $S_n = \sum_{k=1}^n X_k$. Consider the events

$$A = \{S_n = 2 \text{ for some } n > 0\} = \{\text{win } 2\}$$
$$B = \{S_n = 1 \text{ for some } n > 0\} = \{\text{win } 1\}, \qquad C = \{S_0 = 0\}.$$

By sequential Bayes,

$$\mathsf{P}(A \mid C) = \mathsf{P}(A \cap B \mid C) = \mathsf{P}(A \mid B \cap C) \mathsf{P}(B \mid C)$$
$$\stackrel{\text{!!}}{=} \mathsf{P}(1 \rightsquigarrow 2) \mathsf{P}(0 \rightsquigarrow 1) \stackrel{\text{!!}}{=} \mathsf{P}(B \mid C)^2.$$

In other words, $P_0(\{win \ 2\}) = P_0(\{win \ 1\})^2$.

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RANDOM VARIABLES

Definition

An integer-valued random variable X is a map $\Omega o \mathbb{Z}$;

its *distribution* is given by

$$p_k = \mathsf{P}(X = k), \qquad k \in \mathbb{Z}$$
;

its expectation EX is the number

$$\mathsf{E} X = \sum_{k \in \mathbb{Z}} k \, \mathsf{p}_k \in [-\infty, +\infty] \, .$$

INDEPENDENCE OF RANDOM VARIABLES

Definition

Random variables X and Y are **independent**, if for **all** A, $B \subseteq \mathbb{R}$,

$$\mathsf{P}(X \in A, Y \in B) = \mathsf{P}(X \in A) \mathsf{P}(Y \in B).$$

Key property:

If X and Y are *independent* rv's in \mathbb{Z} , and f and g are real functions, then f(X) and g(Y) are **independent** rv's.

Example:

If X, Y are independent integer-valued rv's, then

$$\mathsf{E}(s^{X+Y}) = \mathsf{E}(s^X)\mathsf{E}(s^Y) \,.$$

[similarly, for any countable collection of rv's]

CONDITIONAL EXPECTATION

If an event A is such that P(A) > 0, then the *conditional* expectation of a random variable X given A is

$$\mathsf{E}(X \mid A) = \sum_{k \in \mathbb{Z}} k \mathsf{P}(X = k \mid A).$$

Partition theorem for expectations

If events B_1 , B_2 , form a partition of Ω , then

$$\mathsf{E}(X) = \sum_{k=1}^{\infty} \mathsf{E}(X \mid B_k) \cdot \mathsf{P}(B_k),$$

provided all expressions on the right are finite.

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Example:

Let random variables X_1, X_2, \ldots be independent identically distributed, with values in $\{-K, \ldots, K\}$. Put $S_n = \sum_{k=1}^n X_k$.

Let, further, $N \ge 1$ be an integer valued random variable independent of the sequence X_k , $k \ge 1$. Assume that $EN < \infty$ and denote $S_N = X_1 + \cdots + X_N$. Find ES_N .

For every $n \in \mathbb{N}$, we have

$$\mathsf{E}(S_N \mid N = n) = n\mathsf{E}(X_1 \mid N = n) = n\mathsf{E}(X_1)$$

so that

$$\mathsf{E}(S_N) = \sum_{n=1}^{\infty} \mathsf{E}(S_N \mid N = n) \mathsf{P}(N = n)$$

= $\mathsf{E}(X_1) \sum_{n=1}^{\infty} n \mathsf{P}(N = n) = \mathsf{E}(X_1) \mathsf{E}(N) .$

Exercise: Use the same approach to compute $Var S_N$.

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